On the stability of the distribution function of the composed random variables by their index random variable

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Abstract. Let us consider the composed random variable $\eta = \sum_{k=1}^{\nu} \xi_k$, where $\xi_1, \xi_2, ...$ are independent identically distributed random variables and ν is a positive value random, independent of all ξ_k .

In [1] and [2], we gave some the stabilities of the distribution function of η in the following sense: the small changes in the distribution function of ξ_k only lead to the small changes in the distribution function of η .

In the paper, we investigate the distribution function of η when we have the small changes of the distribution of ν .

1. Introduction

Let us consider the random variable (r.v):

$$\eta = \sum_{k=1}^{\nu} \xi_k \tag{1}$$

where $\xi_1, \xi_2, ...$ are independent identically distributed random variables with the distribution function F(x), ν is a positive value r.v independent of all ξ_k and ν has the distribution function A(x).

In [1] and [2], η is called to be the composed r.v and ν is called to be its index r.v. If $\Psi(x)$ is the distribution function of η with the characteristic function $\psi(x)$ respectively then (see [1] or [2])

$$\psi(x) = a[\varphi(t)] \tag{2}$$

where a(z) is the generating function of ν and $\varphi(t)$ is the characteristic function of ξ_k .

In [1] and [2], we gave some the stabilities of $\Psi(x)$ in the following sence: the small changes in the distribution function F(x) only lead to the small changes in the distribution function $\Psi(x)$.

In this paper, we shall investigate the stability of η 's distribution function when we have the small change of the distribution of the index r.v ν .

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2. Stability theorem

Let us consider the r.v now:

$$\eta_1 = \sum_{k=1}^{\nu_1} \xi_k \tag{3}$$

where ν_1 has the distribution function $A_1(x)$ with the generating function $a_1(z)$. Suppose ξ_k have the stable law with the characteristic function

$$\varphi(t) = \exp\{i\mu t - c|t|^{\alpha}[1 - e\beta \frac{t}{|t|}\omega(t;\alpha)]\}$$
(4)

where c, μ, α, β are real number, $c \ge 0$; $|\beta| \le 1$,

$$2 \ge \alpha \ge \alpha_1 > 1; \quad \omega(t;\alpha) = tg \frac{\alpha t}{2}.$$
 (5)

For every $\varepsilon > 0$ is given, such that

$$\varepsilon < (\frac{\pi}{3c_2})^3 \tag{6}$$

where $c_2 = (c + c|\beta||tg\frac{\alpha_1\pi}{2} + |\mu|)$. We have the following theorem:

Theorem 2.1 (Stability Theorem). Assume that

$$\rho(A; A_1) = \sup_{x \in R^1} |A(x) - A_1(x)| \le \varepsilon$$

$$\mu_A^{\alpha} = \int_0^{+\infty} z^{\alpha} dA(z) < +\infty; \quad \mu_{A_1}^{\alpha} = \int_0^{+\infty} z^{\alpha} dA_1(z) < +\infty, \quad \forall \alpha > 0.$$
(7)

Then we have

$$\rho(\Psi, \Psi_1) \leqslant K_1 \varepsilon^{1/6}$$

where K_1 is a constant independent of ε , $\Psi(x)$ and $\Psi_1(x)$ are the distribution function of η and η_1 respectively.

Lemma 2.1. Let a is a complex number, $a = \rho e^{i\theta}$, such that $|\theta| \leq \frac{\pi}{3}$; $0 \leq \rho \leq 1$. Then we have the following estimation:

$$|a^{t} - 1| \leq \frac{\sqrt{14}|a - 1|}{(1 - |a - 1|)} \quad (forevery \ t > 0)$$
(8)

Proof. Since $a = \rho(\cos \theta + i \sin \theta)$, it follows that $a^t = \rho^t(\cos t\theta + i \sin t\theta)$. Hence

$$|a^{t} - 1|^{2} = (\rho^{t} \cos t\theta - 1)^{2} + (\rho^{t} \sin t\theta)^{2},$$
(9)

we also have

 $(\rho^t \cos t\theta - 1) = (\rho^t - 1) \cos t\theta + (\cos t\theta - 1),$

Notice that $|1 - \cos x| \leq |x|$ for all x, thus

$$|\rho^t \cos t\theta - 1| \leq |\rho^t - 1| + |t\theta|.$$

On the other hand, since $|\sin u| \leq |u|$ for all u,

$$|a^{t} - 1|^{2} \leq 2|\rho^{t} - 1|^{2} + 2t^{2}\theta^{2} + \rho^{2t}t^{2}\theta^{2}, \qquad (10)$$

we can see

$$|a - 1|^{2} = (\rho \cos \theta - 1)^{2} + (\rho^{2} \sin^{2} \theta).$$

$$|\rho \sin \theta| \leq |a - 1|. \tag{11}$$

It follows that

Furthermore,

$$|a|-1| \leq |a-1| \Rightarrow |\rho-1| \leq |a-1| \Rightarrow \rho \geq 1 - |a-1|.$$

From (11) we obtain

$$\sin\theta| \le \frac{|a-1|}{\rho} \le \frac{|a-1|}{1-|a-1|}.$$
(12)

(11)

Since
$$|\theta| \leq \frac{\pi}{3} \Rightarrow |\sin \theta| \geq \frac{|\theta|}{2}$$
, so that
 $|\theta| \leq \frac{2|a-1|}{(1-|a-1|)}$. (13)

From (10) and (13), we have

$$|a^{t} - 1|^{2} \leq 2|\rho^{t} - 1|^{2} + \frac{8t^{2}|a - 1|^{2}}{(1 - |a - 1|)^{2}} + 4\frac{\rho^{2t}t^{2}|a - 1|^{2}}{(1 - |a - 1|)^{2}}.$$
(14)

For all $t \ge 0$, the following inequality holds:

$$1 - \rho^t \leqslant \frac{t(1-\rho)}{\rho}.\tag{15}$$

Using (11) and notice that $|1 - \rho| = |1 - |a|| \le |a - 1|$, we shall have

$$1 - \rho^t \leqslant \frac{t|a-1|}{\rho}.$$
(16)

Hence by (14) we get

$$|a^{t} - 1|^{2} \leq \frac{14t^{2}|a - 1|^{2}}{(1 - |a - 1|)^{2}}$$

Lemma 2.2. Under the notation in (2), let $\delta(\varepsilon)$ be sufficiently small postive number such that $\delta(\varepsilon) \to 0$ when $\varepsilon \to 0$ and

$$|arg\varphi(t)| \leqslant rac{\pi}{3} \quad orall t, \quad |t| \leqslant \delta(arepsilon).$$

Then

$$|\psi(t) - \psi_1(t)| \leq C|t| \quad \forall t, \quad |t| \leq \delta(\varepsilon)$$

where C is a constant independent of ε and $\psi_1(t)$ is the characteristic function with the distribution function $\Psi_1(t)$ respectively.

Proof. We have

$$|\psi(t) - \psi_1(t)| = |\int_0^{+\infty} |\varphi(t)|^z d[A(z) - A_1(z)]| \leq \int_0^{+\infty} |\varphi^z(t) - 1| d[A(z) + A_1(z)].$$
(17)

Notice that, for all $t \in R^1$

$$|e^{itx} - 1| \leq 3|\sin(\frac{tx}{2})| \leq \frac{3}{2}|tx| < 2|tx|.$$

Hence, if we put

$$\mu_F = \int_{-\infty}^{+\infty} |x| dF(x) < +\infty; \quad \varphi(t) = \int_{-\infty}^{+\infty} e^{itx} dF(x),$$

then

$$\varphi(t) - 1| \leq \int |e^{itx} - 1| dF < 2|t|\mu_F.$$

From lemma 2.1, (with $a = \varphi(t); |t| \leq \delta(\varepsilon)$)

$$|\varphi(t) - 1| \leq \frac{\sqrt{14}z|\varphi(t) - 1|}{(1 - |\varphi(t) - 1|)}.$$
(18)

Because there exits moments (from (7)) and with t, $|t| \leq \delta(\varepsilon)$ we can see $|1 - \varphi(t)| \leq \frac{1}{2}$, therefore

$$|\psi(t) - \psi_1(t)| \leq \int_0^{+\infty} \frac{\sqrt{14}z|\varphi(t) - 1|}{(1 - |\varphi(t) - 1|)} d[A(z) + A_1(z)] \leq 4\sqrt{14}\mu_F(\mu_A + \mu_{A_1})|t| = C|t|$$

(do $|\varphi(t) - 1| \leq \mu_F |t| \quad \forall t$) where C is a constant independent of ε and $\mu_F = \int_{-\infty}^{+\infty} |x| dF(x) < \infty$. Proof of Theorem 2.1.

For every N > 0 and $t \in R^1$, we have

$$\begin{aligned} |\psi(t) - \psi_{1}(t)| &= |\int_{0}^{+\infty} \varphi^{z}(t) d[A(z) - A_{1}(z)]| \\ &\leqslant |\int_{0}^{N} \varphi^{z}(t) d[A(z) - A_{1}(z)]| + |\int_{N}^{+\infty} \varphi^{z}(t) d[A(z) - A_{1}(z)]| \\ &\leqslant |[A(z) - A_{1}(z)]|_{0}^{N}| + \int_{0}^{N} |A(z) - A_{1}(z)||\varphi^{z}(t)|| \ln \varphi(t)|dz + \int_{N}^{+\infty} d[A(z) + A_{1}(z)] \\ &= I_{1} + I_{2} + I_{3}. \end{aligned}$$
(19)

First, it easy to see that

$$I_1 \leqslant 2\varepsilon.$$
 (20)

In order to estimate I_2 , notice that $\varphi(t)$ has form (4) with the condition (5) so we have

$$\ln \varphi(t) \leq |\mu| |t| + |t|^{\alpha} (c + c|\beta| |tg\frac{\alpha \pi}{2}|) \leq |\mu| |t| + C_1 |t|^{\alpha}$$
(21)

where $C_1 = c + c|\beta||tg\frac{\alpha\pi}{2}| \leq c + c|\beta||tg\frac{\alpha_1\pi}{2}|$. If $T = T(\varepsilon)$ is a positive number which will be chosen later $(T(\varepsilon) \to \infty \text{ when } \varepsilon \to 0)$, we can see that

$$\begin{split} |\ln \varphi(t)| &\leq |\mu|T + C_1 T^{\alpha} \leq (C_1 + |\mu|) T^{\alpha} \leq C_2 T^{\alpha} \quad \forall t, |t| \leq T(\varepsilon) \\ \text{where } C_2 &= c + c|\beta| |tg \frac{\alpha_1 \pi}{2}| + |\mu|; \quad (\alpha \geq \alpha_1 > 1). \\ \text{Then} \\ \end{split}$$

$$I_2 \leqslant \varepsilon \int_0^N C_2 T^\alpha dz \leqslant C_2 \varepsilon T^\alpha N.$$
(22)

Finally, with α from condition (5), we have

$$I_3 \leqslant \frac{\mu_A^{\alpha} + \mu_{A_1}^{\alpha}}{N^{\alpha}}.$$
(23)

By using (19), (20), (21), (23), we conclude that

$$\psi(t) - \psi_1(t) | \leq 2\varepsilon + C_2 \varepsilon T^{\alpha} N + \frac{\mu_A^{\alpha} + \mu_{A_1}^{\alpha}}{N^{\alpha}}.$$
(24)

Choosing $T = \varepsilon^{-\frac{1}{3^{\alpha}}}$ and $N = T = \varepsilon^{-\frac{1}{3^{\alpha}}}$, we can see that

$$C_{2}\varepsilon T^{\alpha}N \leqslant C_{2}\varepsilon^{1-\frac{1}{3}-\frac{1}{3}} = C_{2}\varepsilon^{\frac{1}{3}},$$
$$(\mu_{A}^{\alpha}+\mu_{A_{1}}^{\alpha})N^{-\alpha} = (\mu_{A}^{\alpha}+\mu_{A_{1}}^{\alpha})\varepsilon^{\frac{1}{3}}.$$

Thus

$$|\psi(t) - \psi_1(t)| \leq 2\varepsilon + C_2 \varepsilon^{\frac{1}{3}} + (\mu_a^{\alpha} + \mu_{A_1}^{\alpha})\varepsilon^{\frac{1}{3}} = C_3 \varepsilon^{\frac{1}{3}}$$
(25)

for every t with $|t| \leq T = \varepsilon^{-3\alpha}$ and C_3 is a constant independent of ε .

For all $\delta(\varepsilon) > 0$, we consider now

$$\int_{-T}^{T} |\frac{\varphi(t) - \varphi_1(t)}{t}| dt = \int_{-\delta(\varepsilon)}^{\delta(\varepsilon)} |\frac{\varphi(t) - \varphi_1(t)}{t}| dt + \int_{\delta(\varepsilon) \le |t| \le T} |\frac{\varphi(t) - \varphi_1(t)}{t}| dt$$

Since

 $\ln z = \ln |z| + i \arg(z) \quad (0 \leqslant \arg z \leqslant 2\pi),$

for all complex number z, letting $z = \varphi(t)$, $(|t| \leq \delta(\varepsilon))$

$$|arg\varphi(t)| \leq |\ln \varphi(t)| \leq C_2 \delta(\varepsilon)$$

with
$$\delta(\varepsilon) = \varepsilon^{\frac{1}{3}}$$
, we shall get $|arg\varphi(t)| \leq C_2 \varepsilon^{\frac{1}{3}}$ and from (6)
 $C_2 \varepsilon^{\frac{1}{3}} \leq \frac{\pi}{3} \Rightarrow |arg\varphi(t)| \leq \frac{\pi}{3}$ for every $t, |t| \leq \delta \varepsilon$.

Hence, using lemma 2.2, we obtain:

$$\int_{-\delta(\varepsilon)}^{\delta(\varepsilon)} \left| \frac{\varphi(t) - \varphi_1(t)}{t} \right| dt \leqslant 2C\delta(\varepsilon) = 2C\varepsilon^{\frac{1}{3}}.$$
(26)

On the other hand, using (25), we get

$$\int_{\delta(\varepsilon) \le |t| \le T} \left| \frac{\varphi(t) - \varphi_1(t)}{t} \right| dt \le C_3 \varepsilon^{\frac{1}{3}} \int_{\delta(\varepsilon)}^T \frac{dt}{t} = C_3 \varepsilon^{\frac{1}{3}} \ln \frac{T}{\delta(\varepsilon)} = C_3 \varepsilon^{\frac{1}{3}} \ln \left(\frac{1}{\frac{1+\alpha}{\varepsilon}}\right) \le C_4 \varepsilon^{\frac{1}{6}}.$$
 (27)

From (26) and (27)

$$\int_{-T}^{T} |\frac{\varphi(t) - \varphi_1(t)}{t}| dt \leq 2C\varepsilon^{\frac{1}{3}} + C_4\varepsilon^{\frac{1}{6}} \leq C_5\varepsilon^{\frac{1}{6}}$$

where C_5 is constant independent of ε .

Indeed, by using Essen's inequality (see [3]) we have

$$\rho(\Psi;\Psi_1) \leqslant C_5 \varepsilon \overline{\overrightarrow{6}} + C_6 \varepsilon \overline{\overrightarrow{4}} \leqslant K_1 \varepsilon \overline{\overrightarrow{6}}$$

where K_1 is a constant independent of ε .

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References

- [1] Tran Kim Thanh, On the characterization of the distribution of the composed random variables and their stabilities Doctor thesis, Hanoi 2000.
- [2] Tran Kim Thanh, Nguyen Huu Bao, On the geometric composed variables and the estimate of the stable degree of the Renyi's characteristic theorem, Acta Mathemaica Vietnamica 21 (1996) 269.
- [3] C. G. Essen, Fourier analysis of distribution functions, Acta Math. 77 (1945) 125.