

ON THE STABILITY OF A CHARACTERIZATION OF EXPONENTIAL DISTRIBUTION BY GEOMETRIC COMPOUNDING

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I. INTRODUCTION

Let X_1, X_2, \dots be nonnegative i.i.d.rv's (independent identically distributed random variables) with $F(x) = P(X_j < x)$, $\varphi(t) = Ee^{itX_j}$, $m = EX_j < +\infty$ and let V be independent of $X_j, j = 1, 2, \dots$ with geometric distribution, i.e.,

$$P(V = k) = p \cdot q^{k-1}, \quad k = 1, 2, \dots, (0 < p < 1, q = 1 - p).$$

The random variable $Z = X_1 + X_2 + \dots + X_V$ is called geometric composed variable of X_j 's.

The notation $G_p(x)$ will mean $P(pZ < x)$; the notation $\varphi_{pZ}(t)$ will mean Ee^{itpZ} . $F_0(x)$ and $\varphi_0(t)$ will denote distribution function and characteristic function respectively of the exponential distribution. In [4], Renyi characterized the exponential distribution proving the following assertions:

- (i) $\lim_{p \rightarrow 0} \bar{G}_p(x) = e^{-x}$; ($\bar{G}_p(x) = P(p \cdot Z \geq x)$),
- (ii) $\bar{G}_p(x) = \bar{F}(x) \Leftrightarrow F(x) = e^{-x}$; ($\bar{F}(x) = 1 - F(x)$).

In [2], we estimated the stable degree of this characterization in the case of the distribution function $F(x)$ being ϵ -exponential, i.e., $\exists T(\epsilon) > 0, T(\epsilon) \rightarrow 0$ when $\epsilon \rightarrow 0$, such that

$$|\varphi(t) - \varphi_0(t)| \leq \epsilon, \forall t : |t| \leq T(\epsilon) \tag{1a}$$

with metrics

$$\lambda(F_1, F_2) = \min_{T > 0} \max \left\{ \max_{|t| \leq T} \frac{1}{2} |\varphi_1(t) - \varphi_2(t)|; \frac{1}{T} \right\},$$

$$\rho(F_1, F_2) = \sup_x |F_1(x) - F_2(x)|$$

for $F_1(x), F_2(x)$ are two distribution function and $\varphi_1(t), \varphi_2(t)$ are characteristic functions respectively.

In this paper we shall consider the stability of Renyi's characterization in the case of the distribution function $G_p(x)$ being ϵ -exponential, i.e., $\exists T(\epsilon) > 0, T(\epsilon) \rightarrow 0$ when $\epsilon \rightarrow 0$, such that

$$|\varphi_{pZ}(t) - \varphi_0(t)| \leq \epsilon \quad \forall t : |t| \leq T(\epsilon), \quad (1b)$$

with the uniform metric ρ .

2. STABILITY THEOREMS

Lemma.

a) Z and $p.Z$ have the finite moment of first degree

$$EZ = \frac{m}{p}, \quad E(pZ) = m. \quad (1)$$

b)

$$\varphi_{pZ}(t) = \frac{p \cdot \varphi(p \cdot t)}{1 - p \cdot \varphi(p \cdot t)}, \quad t \in R. \quad (2)$$

c)

$$|\varphi(t) - 1| \leq m \cdot |t|, \quad \forall t \in R. \quad (3)$$

Proof:

a) We have:

$$\begin{aligned} EZ &= E[E(Z|V)] = E\left[\sum_{k=1}^{\infty} E(Z|V=k) I_{\{V=k\}}(\omega)\right] \\ &= \sum_{k=1}^{\infty} E[X_1 + \dots + X_k] \cdot p \cdot q^{k-1} = \sum_{k=1}^{\infty} m \cdot k \cdot p \cdot q^{k-1} = mp \sum_{k=1}^{\infty} k \cdot q^{k-1} \\ &= mp \cdot \frac{d}{dq} \left(\frac{1}{1-q} \right) = mp \frac{1}{(1-q)^2} = \frac{m}{p}. \end{aligned}$$

Hence, we get (1).

b) By a similar way, it follows that

$$\begin{aligned} \varphi_Z(t) &= Ee^{itZ} = E[E(e^{itZ}|V)] = E\left[\sum_{k=1}^{\infty} E(e^{itZ}|V=k) I_{\{V=k\}}(\omega)\right] \\ &= \sum_{k=1}^{\infty} Ee^{it(X_1 + \dots + X_k)} p \cdot q^{k-1} = \sum_{k=1}^{\infty} \varphi^k(t) \cdot p \cdot q^{k-1} = p\varphi(t) \cdot \sum_{k=0}^{\infty} [q \cdot \varphi(t)]^k = \frac{p\varphi(t)}{1 - q\varphi(t)}. \end{aligned}$$

Hence:

$$\varphi_{pZ}(t) = \varphi_Z(pt) = \frac{p\varphi(pt)}{1 - q\varphi(pt)}.$$

So, we get (2).

c) It is easily seen that

$$|e^{i\alpha} - 1| \leq \alpha, \forall \alpha \in R. \quad (4)$$

Therefore, we get the following estimation:

$$\begin{aligned} |\varphi(t) - 1| &= \left| \int_{-\infty}^{+\infty} (e^{itx} - 1) dF(x) \right| \leq \int_{-\infty}^{+\infty} |e^{itx} - 1| dF(x) \\ &\leq |t| \int_{-\infty}^{+\infty} |x| dF(x) = m. |t|, \quad \forall t \in R, \end{aligned}$$

(where we used (4) with $\alpha = t.x$.) This completes the proof of the Lemma. \diamond

Theorem 1. If $|\varphi_{pZ}(t) - \varphi_0(t)| < \epsilon, \forall t: |t| \leq T$ (for $0 < \epsilon < p/q$ and for some $T > 0$), then we have

$$|\varphi(t) - \varphi_0(t)| < \frac{\epsilon}{p - q.\epsilon}, \quad \forall t: |t| \leq p.T. \quad (5)$$

Proof:

Denote: $r(t) = \varphi_{pZ}(t) - \varphi_0(t)$.

It follows readily from the hypothesis that

$$\begin{aligned} |r(t)| &\leq \epsilon, \quad \forall t: |t| \leq T, \\ |r(t/p)| &\leq \epsilon, \quad \forall t: |t| \leq p.T. \end{aligned} \quad (6)$$

From (2) we have

$$\varphi(u) = \frac{\varphi_{pZ}(u/p)}{p + q\varphi_{pZ}(u/p)}, \quad \forall u \in R. \quad (7)$$

Therefore, from (7) we get the following estimations

$$\begin{aligned} |\varphi(u) - \varphi_0(u)| &= \left| \frac{\varphi_{pZ}(u/p)}{p + q\varphi_{pZ}(u/p)} - \frac{1}{1 - iu} \right| = \left| \frac{\varphi_{pZ}(u/p) - iu\varphi_{pZ}(u/p) - p - q\varphi_{pZ}(u/p)}{(1 - iu)(p + q\varphi_{pZ}(u/p))} \right| \\ &= \left| \frac{(p - iu)\varphi_{pZ}(u/p) - p}{(1 - iu)(p + q\varphi_{pZ}(u/p))} \right| = \left| \frac{[r(u/p) + \varphi_0(u/p)](p - iu) - p}{(1 - iu)(p + q\varphi_{pZ}(u/p))} \right| \\ &= \left| \frac{(p - iu).r(u/p)}{(1 - iu)(p + q\varphi_{pZ}(u/p))} \right| = \frac{|p - iu|}{|1 - iu|} \cdot \frac{|r(u/p)|}{|p + q\varphi_{pZ}(u/p)|} \leq \frac{|r(u/p)|}{p + q\varphi_{pZ}(u/p)}. \end{aligned}$$

Consequently,

$$|\varphi(u) - \varphi_0(u)| \leq \frac{|r(u/p)|}{|p + q\varphi_{pZ}(u/p)|}, \quad \forall u \in R. \quad (8)$$

Notice that $|v| \geq \max\{|\Im v|, |\Re v|\}$ for any complex number $v \in C$. Hence,

$$\begin{aligned}
|p + q\varphi_{pz}(u/p)| &= |p + q[r(u/p) + \varphi_0(u/p)]| = |p + q[r(u/p) + \frac{1}{1 - i.u/p}]| \\
&= \left| p + \frac{pq}{p - iu} + qr(u/p) \right| = \left| p + \frac{qp^2}{p^2 + u^2} + qr(u/p) + iu \frac{pq}{p^2 + u^2} \right| \\
&\geq \left| \Re \left\{ p + \frac{qp^2}{p^2 + u^2} + qr(u/p) + iu \cdot \frac{pq}{p^2 + u^2} \right\} \right| = \left| p + \frac{qp^2}{p^2 + u^2} + q \cdot \Re r(u/p) \right| \\
&\geq p + \frac{qp^2}{p^2 + u^2} - q \cdot |\Re r(u/p)| > p - q |\Re r(u/p)| \geq p - q\epsilon > 0 \quad \forall u : |u| \leq p.T.
\end{aligned}$$

Consequently,

$$|p + q\varphi_{pz}(u/p)| > p - q\epsilon > 0, \quad \forall u : |u| \leq p.T \quad (9)$$

By (8) and (9) we get

$$|\varphi(u) - \varphi_0(u)| \leq \frac{|r(\frac{u}{p})|}{|p + q\varphi_{pz}(u/p)|} < \frac{\epsilon}{p - q\epsilon}, \quad \forall u : |u| \leq p.T.$$

This completes the proof of Theorem 1. \diamond

Theorem 2. Assume that $G_p(x)$ is ϵ -exponential distribution function with the number $T(\epsilon)$ in (1b) satisfying the condition $T(\epsilon) = O(\epsilon^{-\alpha})$ for some $\alpha > 0$ (when $\epsilon \rightarrow 0$). Then we have

$$\rho(F, F_0) \leq K_1 \cdot \epsilon^\alpha - K_2 \epsilon \cdot \ln \epsilon, \quad (10)$$

where $0 < \epsilon < \min\{1, \frac{p}{q}\}$ and $K_1 > 0, K_2 > 0$ are constant numbers independent of ϵ .

Proof:

At first, since $F_0(x)$ is exponential distribution function then $\sup |F_0'(x)| = 1$. Using Essen's inequality (see [3]) with $T = p.T(\epsilon)$ we get the following estimation:

$$\begin{aligned}
\rho(F, F_0) &\leq \frac{1}{\pi} \int_{-p.T(\epsilon)}^{p.T(\epsilon)} \left| \frac{\varphi(t) - \varphi_0(t)}{t} \right| dt + \frac{24}{\pi p.T(\epsilon)} \cdot \sup_x |F_0'(x)| \\
&= \frac{1}{\pi} \int_{-p.T(\epsilon)}^{p.T(\epsilon)} \left| \frac{\varphi(t) - \varphi_0(t)}{t} \right| dt + \frac{24}{\pi p.T(\epsilon)} \leq \\
&\leq \frac{1}{\pi} \int_{|t| \leq \delta} \left| \frac{\varphi(t) - \varphi_0(t)}{t} \right| dt + \int_{\delta < |t| \leq p.T(\epsilon)} \left| \frac{\varphi(t) - \varphi_0(t)}{t} \right| dt + \frac{24}{\pi p.T(\epsilon)} \\
&\leq \frac{1}{\pi} [J_1 + J_2] + \frac{24}{\pi p.T(\epsilon)}, \quad \forall \delta : 0 < \delta < p.T(\epsilon).
\end{aligned} \quad (11)$$

Where

$$J_1 = \int_{|t| \leq \delta} \left| \frac{\varphi(t) - \varphi_0(t)}{t} \right| dt; \quad J_2 = \int_{\delta < |t| \leq p.T(\epsilon)} \left| \frac{\varphi(t) - \varphi_0(t)}{t} \right| dt.$$

In order to estimate J_1 , we put

$$J_1^* = \int_{|t| \leq \delta} \left| \frac{\varphi(t) - 1}{t} \right| dt; \quad J_2^* = \int_{|t| \leq \delta} \left| \frac{\varphi_0(t) - 1}{t} \right| dt.$$

Using Lemma above, it yields

$$J_1^* \leq m \int_{|t| \leq \delta} dt = 2\delta.m,$$

$$J_2^* \leq m_0 \cdot \int_{|t| \leq \delta} dt = 2\delta.m_0 = 2\delta \quad (m_0 = \int_{-\infty}^{+\infty} x dF_0(x) = 1).$$

Therefore,

$$J_1 \leq J_1^* + J_2^* \leq 2\delta(1 + m). \quad (12)$$

On the other hand, according to Theorem 1, we have

$$J_2 \leq \frac{\epsilon}{p - q\epsilon} \int_{\delta < |t| \leq pT(\epsilon)} \frac{1}{|t|} dt = \frac{2\epsilon}{p - q\epsilon} \int_{\delta}^{p.T(\epsilon)} \frac{dt}{t} = \frac{2\epsilon}{p - q\epsilon} \cdot \ln \frac{p.T(\epsilon)}{\delta}.$$

Choosing $\delta = \epsilon^\beta$ with $\beta > \max\{1, \alpha\}$, we get

$$J_2 \leq \frac{2\epsilon}{p - q\epsilon} \ln \frac{p.T(\epsilon)}{\epsilon^\beta}. \quad (13)$$

By (11), (12) and (13), it follows that

$$\rho(F, F_0) \leq \frac{2}{\pi}(1 + m)\epsilon^\beta + \frac{24}{\pi p.T(\epsilon)} + \frac{2\epsilon}{(p - q\epsilon)\pi} \cdot \ln \frac{p.T(\epsilon)}{\epsilon^\beta}. \quad (14)$$

According to the hypothesis $T(\epsilon) = o(\epsilon^{-\alpha})$ we have $T(\epsilon) = K(\epsilon) \cdot \epsilon^{-\alpha}$, where $K(\epsilon) = \frac{T(\epsilon)}{\epsilon^{-\alpha}} \rightarrow K > 0$ when $\epsilon \rightarrow 0$. Then, for sufficiently small ϵ we have

$$\begin{aligned} \epsilon \ln \frac{p.T(\epsilon)}{\epsilon^\beta} &= \epsilon \ln K(\epsilon) - (\alpha + \beta)\epsilon \ln \epsilon \leq -\epsilon \ln \epsilon - (\alpha + \beta)\epsilon \ln \epsilon \\ &\leq -(\alpha + \beta + 1)\epsilon \ln \epsilon, \end{aligned} \quad (15)$$

$2/\pi(p - q\epsilon) < 3(\pi p)$ (for sufficiently small ϵ).

Since $K(\epsilon) \rightarrow K > 0$ ($\epsilon \rightarrow 0$) then $K(\epsilon) > \frac{K}{2}$ (for sufficiently small ϵ). Consequently,

$$\frac{24}{\pi p.T(\epsilon)} = \frac{24}{\pi p \cdot K(\epsilon)} \cdot \epsilon^\alpha < \frac{48}{\pi p \cdot K} \cdot \epsilon^\alpha. \quad (17)$$

By (14), (15), (16) and (17) we get the following estimation

$$\begin{aligned} \rho(F, F_0) &= \frac{2}{\pi}(1 + m)\epsilon^\alpha + \frac{48}{\pi p \cdot K} \cdot \epsilon^\alpha - \frac{3(\alpha + \beta + 1)}{\pi p} \cdot \epsilon \ln \epsilon \\ &= K_1 \cdot \epsilon^\alpha - K_2 \cdot \epsilon \ln \epsilon, \end{aligned}$$

where $K_1 = \frac{2}{\pi}(1+m) + \frac{48}{\pi p K} > 0$; $K_2 = \frac{3(\alpha+\beta+1)}{\pi p} > 0$.

The proof of Theorem 2 is completed. \diamond

Remarks: The right hand side of (10), i.e the expression $K_1\epsilon^\alpha - K_2\epsilon \cdot \ln \epsilon$

- tends to 0 when $\epsilon \rightarrow 0$,
- is the same degree as ϵ^α , if $0 < \alpha < 1$,
- is the same degree as $\epsilon \ln \epsilon$, if $\alpha \geq 1$.

REFERENCES

- [1] A. Kovats: *On Bivariate Geometric Compounding*. Proceeding of the 5th Pannonian Sump. On Math. Stat. Visegrad, Hungary 1985.
- [2] Tran Kim Thanh and Nguyen Huu Bao: On the Geometric Composed Variable and the Estimate of the Stable Degree of the Renyi's Characteristic Theorem, *Acta Math Vietnamica Volume 21 N° 2(1996)*, pp 269-277.
- [3] A.M. Kagan, Iu.V. Linnik, S.R.Rao. *Characteristic Problems of Mathematical statistics*. Moscow 1972.
- [4] A. Renyi. A Characterization of the Poisson process, *Magyar Tud. Akad. Mat. Kut. Int. Köz.* **1(1976)**, 519-527.

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VỀ TÍNH ỔN ĐỊNH CỦA MỘT ĐẶC TRƯNG PHÂN PHỐI MŨ BỞI VIỆC PHỨC HỢP HÌNH HỌC

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Giả sử X_1, X_2, \dots là các biến ngẫu nhiên không âm, độc lập, cùng phân phối, có kỳ vọng hữu hạn và giả sử V là biến ngẫu nhiên độc lập với tất cả các biến X_1, X_2, \dots và có phân phối hình học tham số p

Biến $Z = X_1 + \dots + X_V$ được gọi là biến phức hợp hình học của các biến X_1, X_2, \dots

Renyi ([4]) đã chỉ ra phân phối mũ của pZ là đặc trưng cho việc các X_j có phân phối mũ.

Trong [2] tính ổn định của kết quả trên đã được xét khi X_1 có phân phối ϵ -mũ với các metric λ và ρ .

Trong bài báo này chúng tôi đã xét tính ổn định của kết quả trên của Renyi khi $p.Z$ có phân phối ϵ -mũ với metric ρ .